Stockholm University
Department of Statistics
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Econometrics I

WRITTEN EXAMINATION
Tuesday November 29, 2016, 10 am - 3 pm

Allowed tools: Pocket calculator

Passing rate: 50% of overall total, which is 100 points. For detailed grading criteria, see the course description.

The exam will be handed back on Monday December 19 at pm in room E487.

For the maximum number of points on each problem detailed and clear solutions are required.

If not indicated otherwise, the disturbance terms u_i in the models are supposed to fulfill the usual requirements of normality, homoscedasticity and independence.

This time there is no Swedish version, but you may answer in Swedish.

1. (32p) One believes that the variable X affects the value of variable Y. There is also possibly a seasonal variation in Y. The following model was fitted using OLS from quarterly data from the years 2000-2005 (24 observations).

$$\hat{Y}_t = 1.31 - 0.73D_{2t} - 0.77D_{3t} - 0.03D_{4t} + 2.34X_t \ ,$$

where D_{jt} is the dummy variable for quarter $j, j = 2, 3, 4$ and $RSS = 7.94$, $ESS = 15.35$.

   (a) Can we reject the hypothesis that none of the variables explain Y? Perform a suitable test at significance level 5%.

   (b) One suspects that the influence of X on Y has changed over time. The following model was then estimated:

$$\hat{Y}_t = 1.16 - 0.27D_{2t} - 0.39D_{3t} + 0.01D_{4t} + 2.03X_t + 0.76D_tX_t$$

   where D_t is a dummy variable for the period 2003-2005 and $RSS = 6.61$.

   Is this second model significantly better than the first one? Use significance level 5%.

(c) Rewrite the first estimated model so that instead you include in the model dummy variables for quarters 1, 2 and 3 and a constant term. If all the dummy variables are put to zero, how would you interpret $\hat{Y}_t$?

(d) If we, for example, excluded X from any of the two models even if we know by some test that it should be included, which negative consequence(s) could that have for the resulting model?

2. (27p) A health economist plans to evaluate whether screening patients on arrival or spending extra money on cleaning is more effective in reducing the incidence of infections by the MRSA bacterium in hospitals. She hypothesizes the following model:

$$MRSA_i = \beta_1 + \beta_2 S_i + \beta_3 C_i + u_i$$

where, in hospital i, $MRSA_i$ is the number of infections per thousand patients, S_i is expenditure per patient on screening, and C_i is expenditure per patient on cleaning. u_i is a disturbance term that satisfies the usual regression model assumptions. In particular, u_i is drawn from a distribution with mean zero and constant variance σ^2. The researcher would like to fit the relationship using a sample of hospitals. Unfortunately, data for individual hospitals are not available. Instead she has to use regional data to fit

$$\overline{MRSA}_j = \beta_1 + \beta_2 \bar{S}_j + \beta_3 \bar{C}_j + \bar{u}_j$$

where $\overline{MRSA}_j$, $\bar{S}_j$, $\bar{C}_j$, $\bar{u}_j$ are the averages of $MRSA_i$,C_i, S_i, u_i for the hospitals in region j. There were different numbers of hospitals in the regions, there being n_j hospitals in region j.

(a) Show that the variance of $\bar{u}_j$ is equal to σ^2/n_j and that a regression using OLS to fit the second equation will be subject to heteroscedasticity.

(b) Assuming that the researcher knows the value of n_j for each region, explain how she could respecify the regression model to make it homoscedastic. State the revised specification and demonstrate mathematically that it is homoscedastic.

(c) Suppose that the researcher did not know the values of n_j . Explain in general terms (not mathematically) how, nevertheless, she could per form t tests relating to the regression coefficients, stating any limitatiions.

3. (18p) Using the same n observations of the variables X and Y, the following models are estimated using OLS:

$$\widehat{logY}_i = \hat{\beta}_1 + \hat{\beta}_2 logX_i$$

$$log(Y_i/X_i) = \hat{\alpha}_1 + \hat{\alpha}_2 logX_i$$

Using the expressions for the estimates, write $\hat{\beta}_2$ in terms of $\hat{\alpha}_2$.

Also, write $\hat{\beta}_1$ in terms of $\hat{\alpha}_1$.

4. (23p)

(a) The Durbin-Watson test and the Runs test are only valid under a normal assumption of the disturbance variables.

(b) When autocorrelation is present, OLS estimates are biased as well as inefficient.

(c) The exclusion of important variables may show up as autocorrelation in the disturbance variables.

(d) In the Logit model, p_i is linear but $log(p_i/(1-p_i))$ is non-linear.

(e) In the model $Y_i = \beta_1 + \beta_2 X_i + u_i$, both $\hat{\beta}_1$ and $\hat{\beta}_2$ are affected if Y_i is multiplied with 5.

(f) In a simple linear regression model R^2 is always the same as $\bar{R}^2$.

(g) If we in the model $Y_i = \beta_1 + \beta_2 x X_i + u_i$ have measurement error in X_i, $Cov(X_i, u_i) \neq 0$.

Formula sheet, Econometrics I, Fall 2016

Under the simple linear model $Y_i = \beta_1 + \beta_2 X_i + u_i$, where $u_i \sim N(0, \sigma^2)$ and given independent pairs of observations $(Y_1, X_1), \ldots, (Y_n, X_n)$, the OLS (and ML) estimators are:

$$\begin{aligned}
\widehat{\beta}_1 &= \overline{Y} - \widehat{\beta}_2 \overline{X} \\
\widehat{\beta}_2 &= \frac{\sum \left(X_i - \overline{X}\right)\left(Y_i - \overline{Y}\right)}{\sum \left(X_i - \overline{X}\right)^2} \\
\widehat{\sigma}^2 &= \frac{RSS}{n-2} = \frac{\sum \left(Y_i - \widehat{Y}_i\right)^2}{n-2}
\end{aligned}$$

where $\widehat{Y}_i = \widehat{\beta}_1 + \widehat{\beta}_2 X_i$ and where $E(\widehat{\beta_1}) = \beta_1$, $E(\widehat{\beta_2}) = \beta_2$ and $E(\widehat{\sigma}^2) = \sigma^2$ and further

$$\begin{aligned}
V(\widehat{\beta}_1) &= \frac{\sum X_i^2}{n \sum \left(X_i - \overline{X}\right)^2} \sigma^2 \\
V(\widehat{\beta}_2) &= \frac{\sigma^2}{\sum \left(X_i - \overline{X}\right)^2} \\
V(\widehat{Y}_0) &= \sigma^2 \left(\frac{1}{n} + \frac{\left(X_0 - \overline{X}\right)^2}{\sum \left(X_i - \overline{X}\right)^2} \right) \\
V(Y_0 - \widehat{Y}_0) &= \sigma^2 \left(1 + \frac{1}{n} + \frac{\left(X_0 - \overline{X}\right)^2}{\sum \left(X_i - \overline{X}\right)^2} \right)
\end{aligned}$$

Distributional results:

$$\begin{aligned}
\frac{\widehat{\beta}_i - \beta_i}{se(\widehat{\beta}_i)} &\sim t(n-2), \quad i = 1, 2 \\
\frac{\widehat{\sigma}^2 (n-2)}{\sigma^2} &\sim \chi^2(n-2)
\end{aligned}$$

Coefficient of determination:

$$r^2 = \frac{ESS}{TSS} = 1 - \frac{RSS}{TSS} = 1 - \frac{\sum \left(Y_i - \widehat{Y}_i\right)^2}{\sum \left(Y_i - \overline{Y}\right)^2}$$

Coefficient of correlation:

$$r = \frac{\sum \left(X_i - \overline{X}\right)\left(Y_i - \overline{Y}\right)}{\sqrt{\sum \left(X_i - \overline{X}\right)^2 \sum \left(Y_i - \overline{Y}\right)^2}}$$

where $r = \pm\sqrt{r^2}$

If we let $Y_i^* = w_1 Y_i$ and $X_i^* = w_2 X_2$, then

$$\widehat{\beta}_1^* = w_1 \widehat{\beta}_1, \quad \widehat{\beta}_2^* = \left(\frac{w_1}{w_2}\right) \widehat{\beta}_2, \quad \widehat{\sigma}^{*2} = w_1^2 \widehat{\sigma}^2$$

Under the multiple linear regression model $Y_i = \beta_1 + \beta_2 X_{2i} + \cdots + \beta_k X_{ki} + u_i$, where $u_i \sim N(0, \sigma^2)$ and given independent vectors of observations $(Y_1, X_{21}, \ldots, X_{k1}), \ldots, (Y_n, X_{2n}, \ldots, X_{kn})$, the following holds for the OLS (ML) estimators:

$$\widehat{\sigma}^2 = \frac{RSS}{n-k} = \frac{\sum \left(Y_i - \widehat{Y}_i\right)^2}{n-k}$$

$$\frac{\widehat{\beta}_i - \beta_i}{se(\widehat{\beta}_i)} \sim t(n-k), \quad i = 1, \ldots, k$$

$$\frac{\widehat{\sigma}^2 (n-k)}{\sigma^2} \sim \chi^2(n-k)$$

The multiple coefficient of determination:

$$R^2 = \frac{ESS}{TSS} = 1 - \frac{RSS}{TSS} = 1 - \frac{\sum \left(Y_i - \widehat{Y}_i\right)^2}{\sum \left(Y_i - \overline{Y}\right)^2}$$

Adjusted:

$$\overline{R}^2 = 1 - \frac{RSS/(n-k)}{TSS/(n-1)}$$

Testing $H_0 : \beta_2 = \ldots \beta_k = 0$:

$$F = \frac{ESS/(k-1)}{RSS/(n-k)} = \frac{\sum \left(\widehat{Y}_i - \overline{Y}\right)^2 /(k-1)}{\sum \left(Y_i - \widehat{Y}_i\right)^2 /(n-k)}$$

Comparing an "old" model with a "new" (larger):

$$\begin{aligned} F &= \frac{(ESS_{new} - ESS_{old})/\text{number of new regressors}}{RSS_{new}/(n - \text{number of parameters in the new model})} \\ &= \frac{(R^2_{new} - R^2_{old})/\text{number of new regressors}}{(1 - R^2_{new})/(n - \text{number of parameters in the new model})} \end{aligned}$$

Comparing an "unrestricted" model with a "restricted":

$$F = \frac{(RSS_R - RSS_{UR})/m}{RSS_{UR}/(n-k)} = \frac{(R^2_{UR} - R^2_R)/m}{(1 - R^2_{UR})/(n-k)}$$

where m is the number of linear constraints and k is the number of parameters in the unrestricted model.

Variance inflation factor:

$$VIF_j = \frac{1}{1 - R_j^2}$$

Auxiliary regression:

$$F_j = \frac{R_j^2/(k-2)}{(1 - R_j^2)/(n-k+1)}$$

where $R_j^2 = R^2$ in the regression of the remaining $(k-2)$ regressors.

White's test for heteroscedasticity:

$$n\,R^2 \overset{\text{appr}}{\sim} \chi^2(df = \text{number of regressors in the auxiliary regression})$$

(Holds under H_0 : no heteroscedasticity.)

For $R =$ number of runs, where $N = N_1 + N_2$ total number of observations:

$$E(R) = \frac{2N_1N_2}{N} + 1$$

$$V(R) = \frac{2N_1N_2(2N_1N_2 - N)}{N^2(N-1)}$$

The Durbin Watson d statistic:

$$d = \frac{\sum_{t=2}^{n} (\widehat{u}_t - \widehat{u}_{t-1})^2}{\sum_{t=1}^{n} \widehat{u}_t^2}$$

Akaike's information criterion:

$$AIC = \frac{e^{2k/n} RSS}{n}$$

Schwartz's information citerion:

$$SIC = \frac{n^{k/n} RSS}{n}$$

Mallow's C_p criterion:

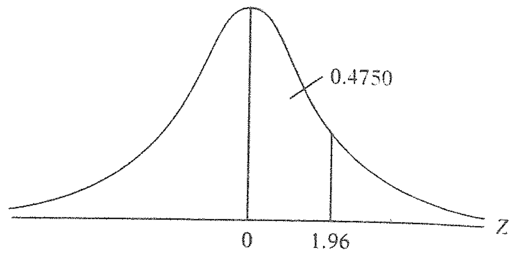
$$C_p = \frac{RSS_p}{\widehat{\sigma}^2} - (n - 2p)$$

TABLE D.1
Areas Under the Standardized Normal Distribution

Example

$\Pr(0 \le Z \le 1.96) = 0.4750$

$\Pr(Z \ge 1.96) = 0.5 - 0.4750 = 0.025$



Z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	.0000	.0040	.0080	.0120	.0160	.0199	.0239	.0279	.0319	.0359
0.1	.0398	.0438	.0478	.0517	.0557	.0596	.0636	.0675	.0714	.0753
0.2	.0793	.0832	.0871	.0910	.0948	.0987	.1026	.1064	.1103	.1141
0.3	.1179	.1217	.1255	.1293	.1331	.1368	.1406	.1443	.1480	.1517
0.4	.1554	.1591	.1628	.1664	.1700	.1736	.1772	.1808	.1844	.1879
0.5	.1915	.1950	.1985	.2019	.2054	.2088	.2123	.2157	.2190	.2224
0.6	.2257	.2291	.2324	.2357	.2389	.2422	.2454	.2486	.2517	.2549
0.7	.2580	.2611	.2642	.2673	.2704	.2734	.2764	.2794	.2823	.2852
0.8	.2881	.2910	.2939	.2967	.2995	.3023	.3051	.3078	.3106	.3133
0.9	.3159	.3186	.3212	.3238	.3264	.3289	.3315	.3340	.3365	.3389
1.0	.3413	.3438	.3461	.3485	.3508	.3531	.3554	.3577	.3599	.3621
1.1	.3643	.3665	.3686	.3708	.3729	.3749	.3770	.3790	.3810	.3830
1.2	.3849	.3869	.3888	.3907	.3925	.3944	.3962	.3980	.3997	.4015
1.3	.4032	.4049	.4066	.4082	.4099	.4115	.4131	.4147	.4162	.4177
1.4	.4192	.4207	.4222	.4236	.4251	.4265	.4279	.4292	.4306	.4319
1.5	.4332	.4345	.4357	.4370	.4382	.4394	.4406	.4418	.4429	.4441
1.6	.4452	.4463	.4474	.4484	.4495	.4505	.4515	.4525	.4535	.4545
1.7	.4454	.4564	.4573	.4582	.4591	.4599	.4608	.4616	.4625	.4633
1.8	.4641	.4649	.4656	.4664	.4671	.4678	.4686	.4693	.4699	.4706
1.9	.4713	.4719	.4726	.4732	.4738	.4744	.4750	.4756	.4761	.4767
2.0	.4772	.4778	.4783	.4788	.4793	.4798	.4803	.4808	.4812	.4817
2.1	.4821	.4826	.4830	.4834	.4838	.4842	.4846	.4850	.4854	.4857
2.2	.4861	.4864	.4868	.4871	.4875	.4878	.4881	.4884	.4887	.4890
2.3	.4893	.4896	.4898	.4901	.4904	.4906	.4909	.4911	.4913	.4916
2.4	.4918	.4920	.4922	.4925	.4927	.4929	.4931	.4932	.4934	.4936
2.5	.4938	.4940	.4941	.4943	.4945	.4946	.4948	.4949	.4951	.4952
2.6	.4953	.4955	.4956	.4957	.4959	.4960	.4961	.4962	.4963	.4964
2.7	.4965	.4966	.4967	.4968	.4969	.4970	.4971	.4972	.4973	.4974
2.8	.4974	.4975	.4976	.4977	.4977	.4978	.4979	.4979	.4980	.4981
2.9	.4981	.4982	.4982	.4983	.4984	.4984	.4985	.4985	.4986	.4986
3.0	.4987	.4987	.4987	.4988	.4988	.4989	.4989	.4989	.4990	.4990

Note: This table gives the area in the right-hand tail of the distribution (i.e., $Z \ge 0$). But since the normal distribution is symmetrical about $Z = 0$, the area in the left-hand tail is the same as the area in the corresponding right-hand tail. For example, $P(-1.96 \le Z \le 0) = 0.4750$. Therefore, $P(-1.96 \le Z \le 1.96) = 2(0.4750) = 0.95$.

TABLE D.2
Percentage Points of the *t* Distribution

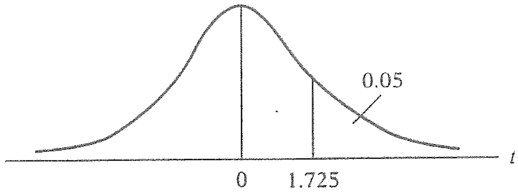
Source: From E. S. Pearson and H. O. Hartley, eds., *Biometrika Tables for Statisticians*, vol. 1, 3d ed., table 12, Cambridge University Press, New York, 1966. Reproduced by permission of the editors and trustees of *Biometrika*.

Example

$\Pr(t > 2.086) = 0.025$

$\Pr(t > 1.725) = 0.05$ for df = 20

$\Pr(|t| > 1.725) = 0.10$



df \ Pr	0.25 0.50	0.10 0.20	0.05 0.10	0.025 0.05	0.01 0.02	0.005 0.010	0.001 0.002
1	1.000	3.078	6.314	12.706	31.821	63.657	318.31
2	0.816	1.886	2.920	4.303	6.965	9.925	22.327
3	0.765	1.638	2.353	3.182	4.541	5.841	10.214
4	0.741	1.533	2.132	2.776	3.747	4.604	7.173
5	0.727	1.476	2.015	2.571	3.365	4.032	5.893
6	0.718	1.440	1.943	2.447	3.143	3.707	5.208
7	0.711	1.415	1.895	2.365	2.998	3.499	4.785
8	0.706	1.397	1.860	2.306	2.896	3.355	4.501
9	0.703	1.383	1.833	2.262	2.821	3.250	4.297
10	0.700	1.372	1.812	2.228	2.764	3.169	4.144
11	0.697	1.363	1.796	2.201	2.718	3.106	4.025
12	0.695	1.356	1.782	2.179	2.681	3.055	3.930
13	0.694	1.350	1.771	2.160	2.650	3.012	3.852
14	0.692	1.345	1.761	2.145	2.624	2.977	3.787
15	0.691	1.341	1.753	2.131	2.602	2.947	3.733
16	0.690	1.337	1.746	2.120	2.583	2.921	3.686
17	0.689	1.333	1.740	2.110	2.567	2.898	3.646
18	0.688	1.330	1.734	2.101	2.552	2.878	3.610
19	0.688	1.328	1.729	2.093	2.539	2.861	3.579
20	0.687	1.325	1.725	2.086	2.528	2.845	3.552
21	0.686	1.323	1.721	2.080	2.518	2.831	3.527
22	0.686	1.321	1.717	2.074	2.508	2.819	3.505
23	0.685	1.319	1.714	2.069	2.500	2.807	3.485
24	0.685	1.318	1.711	2.064	2.492	2.797	3.467
25	0.684	1.316	1.708	2.060	2.485	2.787	3.450
26	0.684	1.315	1.706	2.056	2.479	2.779	3.435
27	0.684	1.314	1.703	2.052	2.473	2.771	3.421
28	0.683	1.313	1.701	2.048	2.467	2.763	3.408
29	0.683	1.311	1.699	2.045	2.462	2.756	3.396
30	0.683	1.310	1.697	2.042	2.457	2.750	3.385
40	0.681	1.303	1.684	2.021	2.423	2.704	3.307
60	0.679	1.296	1.671	2.000	2.390	2.660	3.232
120	0.677	1.289	1.658	1.980	2.358	2.617	3.160
∞	0.674	1.282	1.645	1.960	2.326	2.576	3.090

Note: The smaller probability shown at the head of each column is the area in one tail; the larger probability is the area in both tails.

TABLE D.3 Upper Percentage Points of the F Distribution

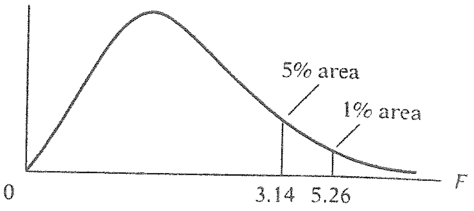
Example

$\Pr(F > 1.59) = 0.25$

$\Pr(F > 2.42) = 0.10$ for df $N_1 = 10$

$\Pr(F > 3.14) = 0.05$ and $N_2 = 9$

$\Pr(F > 5.26) = 0.01$



df for denominator N_2	Pr	df for numerator N_1: 1	2	3	4	5	6	7	8	9	10	11	12
1	.25	5.83	7.50	8.20	8.58	8.82	8.98	9.10	9.19	9.26	9.32	9.36	9.41
	.10	39.9	49.5	53.6	55.8	57.2	58.2	58.9	59.4	59.9	60.2	60.5	60.7
	.05	161	200	216	225	230	234	237	239	241	242	243	244
2	.25	2.57	3.00	3.15	3.23	3.28	3.31	3.34	3.35	3.37	3.38	3.39	3.39
	.10	8.53	9.00	9.16	9.24	9.29	9.33	9.35	9.37	9.38	9.39	9.40	9.41
	.05	18.5	19.0	19.2	19.2	19.3	19.3	19.4	19.4	19.4	19.4	19.4	19.4
	.01	98.5	99.0	99.2	99.2	99.3	99.3	99.4	99.4	99.4	99.4	99.4	99.4
3	.25	2.02	2.28	2.36	2.39	2.41	2.42	2.43	2.44	2.44	2.44	2.45	2.45
	.10	5.54	5.46	5.39	5.34	5.31	5.28	5.27	5.25	5.24	5.23	5.22	5.22
	.05	10.1	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81	8.79	8.76	8.74
	.01	34.1	30.8	29.5	28.7	28.2	27.9	27.7	27.5	27.3	27.2	27.1	27.1
4	.25	1.81	2.00	2.05	2.06	2.07	2.08	2.08	2.08	2.08	2.08	2.08	2.08
	.10	4.54	4.32	4.19	4.11	4.05	4.01	3.98	3.95	3.94	3.92	3.91	3.90
	.05	7.71	6.94	6.59	6.39	6.26	6.16	6.09	6.04	6.00	5.96	5.94	5.91
	.01	21.2	18.0	16.7	16.0	15.5	15.2	15.0	14.8	14.7	14.5	14.4	14.4
5	.25	1.69	1.85	1.88	1.89	1.89	1.89	1.89	1.89	1.89	1.89	1.89	1.89
	.10	4.06	3.78	3.62	3.52	3.45	3.40	3.37	3.34	3.32	3.30	3.28	3.27
	.05	6.61	5.79	5.41	5.19	5.05	4.95	4.88	4.82	4.77	4.74	4.71	4.68
	.01	16.3	13.3	12.1	11.4	11.0	10.7	10.5	10.3	10.2	10.1	9.96	9.89
6	.25	1.62	1.76	1.78	1.79	1.79	1.78	1.78	1.78	1.77	1.77	1.77	1.77
	.10	3.78	3.46	3.29	3.18	3.11	3.05	3.01	2.98	2.96	2.94	2.92	2.90
	.05	5.99	5.14	4.76	4.53	4.39	4.28	4.21	4.15	4.10	4.06	4.03	4.00
	.01	13.7	10.9	9.78	9.15	8.75	8.47	8.26	8.10	7.98	7.87	7.79	7.72
7	.25	1.57	1.70	1.72	1.72	1.71	1.71	1.70	1.70	1.69	1.69	1.69	1.68
	.10	3.59	3.26	3.07	2.96	2.88	2.83	2.78	2.75	2.72	2.70	2.68	2.67
	.05	5.59	4.74	4.35	4.12	3.97	3.87	3.79	3.73	3.68	3.64	3.60	3.57
	.01	12.2	9.55	8.45	7.85	7.46	7.19	6.99	6.84	6.72	6.62	6.54	6.47
8	.25	1.54	1.66	1.67	1.66	1.66	1.65	1.64	1.64	1.63	1.63	1.63	1.62
	.10	3.46	3.11	2.92	2.81	2.73	2.67	2.62	2.59	2.56	2.54	2.52	2.50
	.05	5.32	4.46	4.07	3.84	3.69	3.58	3.50	3.44	3.39	3.35	3.31	3.28
	.01	11.3	8.65	7.59	7.01	6.63	6.37	6.18	6.03	5.91	5.81	5.73	5.67
9	.25	1.51	1.62	1.63	1.63	1.62	1.61	1.60	1.60	1.59	1.59	1.58	1.58
	.10	3.36	3.01	2.81	2.69	2.61	2.55	2.51	2.47	2.44	2.42	2.40	2.38
	.05	5.12	4.26	3.86	3.63	3.48	3.37	3.29	3.23	3.18	3.14	3.10	3.07
	.01	10.6	8.02	6.99	6.42	6.06	5.80	5.61	5.47	5.35	5.26	5.18	5.11

Source: From E. S. Pearson and H. O. Hartley, eds., *Biometrika Tables for Statisticians*, vol. 1, 3d ed., table 18, Cambridge University Press, New York, 1966. Reproduced by permission of the editors and trustees of *Biometrika*.

F-table continued

df for numerator N_1													df for denominator N_2
15	20	24	30	40	50	60	100	120	200	500	∞	Pr	
9.49	9.58	9.63	9.67	9.71	9.74	9.76	9.78	9.80	9.82	9.84	9.85	.25	1
61.2	61.7	62.0	62.3	62.5	62.7	62.8	63.0	63.1	63.2	63.3	63.3	.10	
246	248	249	250	251	252	252	253	253	254	254	254	.05	
3.41	3.43	3.43	3.44	3.45	3.45	3.46	3.47	3.47	3.48	3.48	3.48	.25	2
9.42	9.44	9.45	9.46	9.47	9.47	9.47	9.48	9.48	9.49	9.49	9.49	.10	
19.4	19.4	19.5	19.5	19.5	19.5	19.5	19.5	19.5	19.5	19.5	19.5	.05	
99.4	99.4	99.5	99.5	99.5	99.5	99.5	99.5	99.5	99.5	99.5	99.5	.01	
2.46	2.46	2.46	2.47	2.47	2.47	2.47	2.47	2.47	2.47	2.47	2.47	.25	3
5.20	5.18	5.18	5.17	5.16	5.15	5.15	5.14	5.14	5.14	5.14	5.13	.10	
8.70	8.66	8.64	8.62	8.59	8.58	8.57	8.55	8.55	8.54	8.53	8.53	.05	
26.9	26.7	26.6	26.5	26.4	26.4	26.3	26.2	26.2	26.2	26.1	26.1	.01	
2.08	2.08	2.08	2.08	2.08	2.08	2.08	2.08	2.08	2.08	2.08	2.08	.25	4
3.87	3.84	3.83	3.82	3.80	3.80	3.79	3.78	3.78	3.77	3.76	3.76	.10	
5.86	5.80	5.77	5.75	5.72	5.70	5.69	5.66	5.66	5.65	5.64	5.63	.05	
14.2	14.0	13.9	13.8	13.7	13.7	13.7	13.6	13.6	13.5	13.5	13.5	.01	
1.89	1.88	1.88	1.88	1.88	1.88	1.87	1.87	1.87	1.87	1.87	1.87	.25	5
3.24	3.21	3.19	3.17	3.16	3.15	3.14	3.13	3.12	3.12	3.11	3.10	.10	
4.62	4.56	4.53	4.50	4.46	4.44	4.43	4.41	4.40	4.39	4.37	4.36	.05	
9.72	9.55	9.47	9.38	9.29	9.24	9.20	9.13	9.11	9.08	9.04	9.02	.01	
1.76	1.76	1.75	1.75	1.75	1.75	1.74	1.74	1.74	1.74	1.74	1.74	.25	6
2.87	2.84	2.82	2.80	2.78	2.77	2.76	2.75	2.74	2.73	2.73	2.72	.10	
3.94	3.87	3.84	3.81	3.77	3.75	3.74	3.71	3.70	3.69	3.68	3.67	.05	
7.56	7.40	7.31	7.23	7.14	7.09	7.06	6.99	6.97	6.93	6.90	6.88	.01	
1.68	1.67	1.67	1.66	1.66	1.66	1.65	1.65	1.65	1.65	1.65	1.65	.25	7
2.63	2.59	2.58	2.56	2.54	2.52	2.51	2.50	2.49	2.48	2.48	2.47	.10	
3.51	3.44	3.41	3.38	3.34	3.32	3.30	3.27	3.27	3.25	3.24	3.23	.05	
6.31	6.16	6.07	5.99	5.91	5.86	5.82	5.75	5.74	5.70	5.67	5.65	.01	
1.62	1.61	1.60	1.60	1.59	1.59	1.59	1.58	1.58	1.58	1.58	1.58	.25	8
2.46	2.42	2.40	2.38	2.36	2.35	2.34	2.32	2.32	2.31	2.30	2.29	.10	
3.22	3.15	3.12	3.08	3.04	2.02	3.01	2.97	2.97	2.95	2.94	2.93	.05	
5.52	5.36	5.28	5.20	5.12	5.07	5.03	4.96	4.95	4.91	4.88	4.86	.01	
1.57	1.56	1.56	1.55	1.55	1.54	1.54	1.53	1.53	1.53	1.53	1.53	.25	9
2.34	2.30	2.28	2.25	2.23	2.22	2.21	2.19	2.18	2.17	2.17	2.16	.10	
3.01	2.94	2.90	2.86	2.83	2.80	2.79	2.76	2.75	2.73	2.72	2.71	.05	
4.96	4.81	4.73	4.65	4.57	4.52	4.48	4.42	4.40	4.36	4.33	4.31	.01	

(Continued)

TABLE D.3 Upper Percentage Points of the F Distribution *(Continued)*

df for denominator N_2		df for numerator N_1											
	Pr	1	2	3	4	5	6	7	8	9	10	11	12
10	.25	1.49	1.60	1.60	1.59	1.59	1.58	1.57	1.56	1.56	1.55	1.55	1.54
	.10	3.29	2.92	2.73	2.61	2.52	2.46	2.41	2.38	2.35	2.32	2.30	2.28
	.05	4.96	4.10	3.71	3.48	3.33	3.22	3.14	3.07	3.02	2.98	2.94	2.91
	.01	10.0	7.56	6.55	5.99	5.64	5.39	5.20	5.06	4.94	4.85	4.77	4.71
11	.25	1.47	1.58	1.58	1.57	1.56	1.55	1.54	1.53	1.53	1.52	1.52	1.51
	.10	3.23	2.86	2.66	2.54	2.45	2.39	2.34	2.30	2.27	2.25	2.23	2.21
	.05	4.84	3.98	3.59	3.36	3.20	3.09	3.01	2.95	2.90	2.85	2.82	2.79
	.01	9.65	7.21	6.22	5.67	5.32	5.07	4.89	4.74	4.63	4.54	4.46	4.40
12	.25	1.46	1.56	1.56	1.55	1.54	1.53	1.52	1.51	1.51	1.50	1.50	1.49
	.10	3.18	2.81	2.61	2.48	2.39	2.33	2.28	2.24	2.21	2.19	2.17	2.15
	.05	4.75	3.89	3.49	3.26	3.11	3.00	2.91	2.85	2.80	2.75	2.72	2.69
	.01	9.33	6.93	5.95	5.41	5.06	4.82	4.64	4.50	4.39	4.30	4.22	4.16
13	.25	1.45	1.55	1.55	1.53	1.52	1.51	1.50	1.49	1.49	1.48	1.47	1.47
	.10	3.14	2.76	2.56	2.43	2.35	2.28	2.23	2.20	2.16	2.14	2.12	2.10
	.05	4.67	3.81	3.41	3.18	3.03	2.92	2.83	2.77	2.71	2.67	2.63	2.60
	.01	9.07	6.70	5.74	5.21	4.86	4.62	4.44	4.30	4.19	4.10	4.02	3.96
14	.25	1.44	1.53	1.53	1.52	1.51	1.50	1.49	1.48	1.47	1.46	1.46	1.45
	.10	3.10	2.73	2.52	2.39	2.31	2.24	2.19	2.15	2.12	2.10	2.08	2.05
	.05	4.60	3.74	3.34	3.11	2.96	2.85	2.76	2.70	2.65	2.60	2.57	2.53
	.01	8.86	6.51	5.56	5.04	4.69	4.46	4.28	4.14	4.03	3.94	3.86	3.80
15	.25	1.43	1.52	1.52	1.51	1.49	1.48	1.47	1.46	1.46	1.45	1.44	1.44
	.10	3.07	2.70	2.49	2.36	2.27	2.21	2.16	2.12	2.09	2.06	2.04	2.02
	.05	4.54	3.68	3.29	3.06	2.90	2.79	2.71	2.64	2.59	2.54	2.51	2.48
	.01	8.68	6.36	5.42	4.89	4.56	4.32	4.14	4.00	3.89	3.80	3.73	3.67
16	.25	1.42	1.51	1.51	1.50	1.48	1.47	1.46	1.45	1.44	1.44	1.44	1.43
	.10	3.05	2.67	2.46	2.33	2.24	2.18	2.13	2.09	2.06	2.03	2.01	1.99
	.05	4.49	3.63	3.24	3.01	2.85	2.74	2.66	2.59	2.54	2.49	2.46	2.42
	.01	8.53	6.23	5.29	4.77	4.44	4.20	4.03	3.89	3.78	3.69	3.62	3.55
17	.25	1.42	1.51	1.50	1.49	1.47	1.46	1.45	1.44	1.43	1.43	1.42	1.41
	.10	3.03	2.64	2.44	2.31	2.22	2.15	2.10	2.06	2.03	2.00	1.98	1.96
	.05	4.45	3.59	3.20	2.96	2.81	2.70	2.61	2.55	2.49	2.45	2.41	2.38
	.01	8.40	6.11	5.18	4.67	4.34	4.10	3.93	3.79	3.68	3.59	3.52	3.46
18	.25	1.41	1.50	1.49	1.48	1.46	1.45	1.44	1.43	1.42	1.42	1.41	1.40
	.10	3.01	2.62	2.42	2.29	2.20	2.13	2.08	2.04	2.00	1.98	1.96	1.93
	.05	4.41	3.55	3.16	2.93	2.77	2.66	2.58	2.51	2.46	2.41	2.37	2.34
	.01	8.29	6.01	5.09	4.58	4.25	4.01	3.84	3.71	3.60	3.51	3.43	3.37
19	.25	1.41	1.49	1.49	1.47	1.46	1.44	1.43	1.42	1.41	1.41	1.40	1.40
	.10	2.99	2.61	2.40	2.27	2.18	2.11	2.06	2.02	1.98	1.96	1.94	1.91
	.05	4.38	3.52	3.13	2.90	2.74	2.63	2.54	2.48	2.42	2.38	2.34	2.31
	.01	8.18	5.93	5.01	4.50	4.17	3.94	3.77	3.63	3.52	3.43	3.36	3.30
20	.25	1.40	1.49	1.48	1.46	1.45	1.44	1.43	1.42	1.41	1.40	1.39	1.39
	.10	2.97	2.59	2.38	2.25	2.16	2.09	2.04	2.00	1.96	1.94	1.92	1.89
	.05	4.35	3.49	3.10	2.87	2.71	2.60	2.51	2.45	2.39	2.35	2.31	2.28
	.01	8.10	5.85	4.94	4.43	4.10	3.87	3.70	3.56	3.46	3.37	3.29	3.23

F-table continued

df for numerator N_1: 15	20	24	30	40	50	60	100	120	200	500	∞	Pr	df for denominator N_2
1.53	1.52	1.52	1.51	1.51	1.50	1.50	1.49	1.49	1.49	1.48	1.48	.25	10
2.24	2.20	2.18	2.16	2.13	2.12	2.11	2.09	2.08	2.07	2.06	2.06	.10	
2.85	2.77	2.74	2.70	2.66	2.64	2.62	2.59	2.58	2.56	2.55	2.54	.05	
4.56	4.41	4.33	4.25	4.17	4.12	4.08	4.01	4.00	3.96	3.93	3.91	.01	
1.50	1.49	1.49	1.48	1.47	1.47	1.47	1.46	1.46	1.46	1.45	1.45	.25	11
2.17	2.12	2.10	2.08	2.05	2.04	2.03	2.00	2.00	1.99	1.98	1.97	.10	
2.72	2.65	2.61	2.57	2.53	2.51	2.49	2.46	2.45	2.43	2.42	2.40	.05	
4.25	4.10	4.02	3.94	3.86	3.81	3.78	3.71	3.69	3.66	3.62	3.60	.01	
1.48	1.47	1.46	1.45	1.45	1.44	1.44	1.43	1.43	1.43	1.42	1.42	.25	12
2.10	2.06	2.04	2.01	1.99	1.97	1.96	1.94	1.93	1.92	1.91	1.90	.10	
2.62	2.54	2.51	2.47	2.43	2.40	2.38	2.35	2.34	2.32	2.31	2.30	.05	
4.01	3.86	3.78	3.70	3.62	3.57	3.54	3.47	3.45	3.41	3.38	3.36	.01	
1.46	1.45	1.44	1.43	1.42	1.42	1.42	1.41	1.41	1.40	1.40	1.40	.25	13
2.05	2.01	1.98	1.96	1.93	1.92	1.90	1.88	1.88	1.86	1.85	1.85	.10	
2.53	2.46	2.42	2.38	2.34	2.31	2.30	2.26	2.25	2.23	2.22	2.21	.05	
3.82	3.66	3.59	3.51	3.43	3.38	3.34	3.27	3.25	3.22	3.19	3.17	.01	
1.44	1.43	1.42	1.41	1.41	1.40	1.40	1.39	1.39	1.39	1.38	1.38	.25	14
2.01	1.96	1.94	1.91	1.89	1.87	1.86	1.83	1.83	1.82	1.80	1.80	.10	
2.46	2.39	2.35	2.31	2.27	2.24	2.22	2.19	2.18	2.16	2.14	2.13	.05	
3.66	3.51	3.43	3.35	3.27	3.22	3.18	3.11	3.09	3.06	3.03	3.00	.01	
1.43	1.41	1.41	1.40	1.39	1.39	1.38	1.38	1.37	1.37	1.36	1.36	.25	15
1.97	1.92	1.90	1.87	1.85	1.83	1.82	1.79	1.79	1.77	1.76	1.76	.10	
2.40	2.33	2.29	2.25	2.20	2.18	2.16	2.12	2.11	2.10	2.08	2.07	.05	
3.52	3.37	3.29	3.21	3.13	3.08	3.05	2.98	2.96	2.92	2.89	2.87	.01	
1.41	1.40	1.39	1.38	1.37	1.37	1.36	1.36	1.35	1.35	1.34	1.34	.25	16
1.94	1.89	1.87	1.84	1.81	1.79	1.78	1.76	1.75	1.74	1.73	1.72	.10	
2.35	2.28	2.24	2.19	2.15	2.12	2.11	2.07	2.06	2.04	2.02	2.01	.05	
3.41	3.26	3.18	3.10	3.02	2.97	2.93	2.86	2.84	2.81	2.78	2.75	.01	
1.40	1.39	1.38	1.37	1.36	1.35	1.35	1.34	1.34	1.34	1.33	1.33	.25	17
1.91	1.86	1.84	1.81	1.78	1.76	1.75	1.73	1.72	1.71	1.69	1.69	.10	
2.31	2.23	2.19	2.15	2.10	2.08	2.06	2.02	2.01	1.99	1.97	1.96	.05	
3.31	3.16	3.08	3.00	2.92	2.87	2.83	2.76	2.75	2.71	2.68	2.65	.01	
1.39	1.38	1.37	1.36	1.35	1.34	1.34	1.33	1.33	1.32	1.32	1.32	.25	18
1.89	1.84	1.81	1.78	1.75	1.74	1.72	1.70	1.69	1.68	1.67	1.66	.10	
2.27	2.19	2.15	2.11	2.06	2.04	2.02	1.98	1.97	1.95	1.93	1.92	.05	
3.23	3.08	3.00	2.92	2.84	2.78	2.75	2.68	2.66	2.62	2.59	2.57	.01	
1.38	1.37	1.36	1.35	1.34	1.33	1.33	1.32	1.32	1.31	1.31	1.30	.25	19
1.86	1.81	1.79	1.76	1.73	1.71	1.70	1.67	1.67	1.65	1.64	1.63	.10	
2.23	2.16	2.11	2.07	2.03	2.00	1.98	1.94	1.93	1.91	1.89	1.88	.05	
3.15	3.00	2.92	2.84	2.76	2.71	2.67	2.60	2.58	2.55	2.51	2.49	.01	
1.37	1.36	1.35	1.34	1.33	1.33	1.32	1.31	1.31	1.30	1.30	1.29	.25	20
1.84	1.79	1.77	1.74	1.71	1.69	1.68	1.65	1.64	1.63	1.62	1.61	.10	
2.20	2.12	2.08	2.04	1.99	1.97	1.95	1.91	1.90	1.88	1.86	1.84	.05	
3.09	2.94	2.86	2.78	2.69	2.64	2.61	2.54	2.52	2.48	2.44	2.42	.01	

(Continued

TABLE D.3 **Upper Percentage Points of the F Distribution** *(Continued)*

df for denominator N_2	Pr	df for numerator N_1: 1	2	3	4	5	6	7	8	9	10	11	12
22	.25	1.40	1.48	1.47	1.45	1.44	1.42	1.41	1.40	1.39	1.39	1.38	1.37
	.10	2.95	2.56	2.35	2.22	2.13	2.06	2.01	1.97	1.93	1.90	1.88	1.86
	.05	4.30	3.44	3.05	2.82	2.66	2.55	2.46	2.40	2.34	2.30	2.26	2.23
	.01	7.95	5.72	4.82	4.31	3.99	3.76	3.59	3.45	3.35	3.26	3.18	3.12
24	.25	1.39	1.47	1.46	1.44	1.43	1.41	1.40	1.39	1.38	1.38	1.37	1.36
	.10	2.93	2.54	2.33	2.19	2.10	2.04	1.98	1.94	1.91	1.88	1.85	1.83
	.05	4.26	3.40	3.01	2.78	2.62	2.51	2.42	2.36	2.30	2.25	2.21	2.18
	.01	7.82	5.61	4.72	4.22	3.90	3.67	3.50	3.36	3.26	3.17	3.09	3.03
26	.25	1.38	1.46	1.45	1.44	1.42	1.41	1.39	1.38	1.37	1.37	1.36	1.35
	.10	2.91	2.52	2.31	2.17	2.08	2.01	1.96	1.92	1.88	1.86	1.84	1.81
	.05	4.23	3.37	2.98	2.74	2.59	2.47	2.39	2.32	2.27	2.22	2.18	2.15
	.01	7.72	5.53	4.64	4.14	3.82	3.59	3.42	3.29	3.18	3.09	3.02	2.96
28	.25	1.38	1.46	1.45	1.43	1.41	1.40	1.39	1.38	1.37	1.36	1.35	1.34
	.10	2.89	2.50	2.29	2.16	2.06	2.00	1.94	1.90	1.87	1.84	1.81	1.79
	.05	4.20	3.34	2.95	2.71	2.56	2.45	2.36	2.29	2.24	2.19	2.15	2.12
	.01	7.64	5.45	4.57	4.07	3.75	3.53	3.36	3.23	3.12	3.03	2.96	2.90
30	.25	1.38	1.45	1.44	1.42	1.41	1.39	1.38	1.37	1.36	1.35	1.35	1.34
	.10	2.88	2.49	2.28	2.14	2.05	1.98	1.93	1.88	1.85	1.82	1.79	1.77
	.05	4.17	3.32	2.92	2.69	2.53	2.42	2.33	2.27	2.21	2.16	2.13	2.09
	.01	7.56	5.39	4.51	4.02	3.70	3.47	3.30	3.17	3.07	2.98	2.91	2.84
40	.25	1.36	1.44	1.42	1.40	1.39	1.37	1.36	1.35	1.34	1.33	1.32	1.31
	.10	2.84	2.44	2.23	2.09	2.00	1.93	1.87	1.83	1.79	1.76	1.73	1.71
	.05	4.08	3.23	2.84	2.61	2.45	2.34	2.25	2.18	2.12	2.08	2.04	2.00
	.01	7.31	5.18	4.31	3.83	3.51	3.29	3.12	2.99	2.89	2.80	2.73	2.66
60	.25	1.35	1.42	1.41	1.38	1.37	1.35	1.33	1.32	1.31	1.30	1.29	1.29
	.10	2.79	2.39	2.18	2.04	1.95	1.87	1.82	1.77	1.74	1.71	1.68	1.66
	.05	4.00	3.15	2.76	2.53	2.37	2.25	2.17	2.10	2.04	1.99	1.95	1.92
	.01	7.08	4.98	4.13	3.65	3.34	3.12	2.95	2.82	2.72	2.63	2.56	2.50
120	.25	1.34	1.40	1.39	1.37	1.35	1.33	1.31	1.30	1.29	1.28	1.27	1.26
	.10	2.75	2.35	2.13	1.99	1.90	1.82	1.77	1.72	1.68	1.65	1.62	1.60
	.05	3.92	3.07	2.68	2.45	2.29	2.17	2.09	2.02	1.96	1.91	1.87	1.83
	.01	6.85	4.79	3.95	3.48	3.17	2.96	2.79	2.66	2.56	2.47	2.40	2.34
200	.25	1.33	1.39	1.38	1.36	1.34	1.32	1.31	1.29	1.28	1.27	1.26	1.25
	.10	2.73	2.33	2.11	1.97	1.88	1.80	1.75	1.70	1.66	1.63	1.60	1.57
	.05	3.89	3.04	2.65	2.42	2.26	2.14	2.06	1.98	1.93	1.88	1.84	1.80
	.01	6.76	4.71	3.88	3.41	3.11	2.89	2.73	2.60	2.50	2.41	2.34	2.27
∞	.25	1.32	1.39	1.37	1.35	1.33	1.31	1.29	1.28	1.27	1.25	1.24	1.24
	.10	2.71	2.30	2.08	1.94	1.85	1.77	1.72	1.67	1.63	1.60	1.57	1.55
	.05	3.84	3.00	2.60	2.37	2.21	2.10	2.01	1.94	1.88	1.83	1.79	1.75
	.01	6.63	4.61	3.78	3.32	3.02	2.80	2.64	2.51	2.41	2.32	2.25	2.18

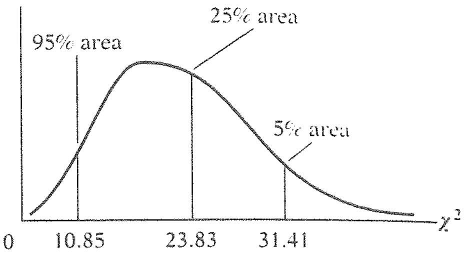
F-table continued

df for numerator N_1: 15	20	24	30	40	50	60	100	120	200	500	∞	Pr	df for denominator N_2
1.36	1.34	1.33	1.32	1.31	1.31	1.30	1.30	1.30	1.29	1.29	1.28	.25	
1.81	1.76	1.73	1.70	1.67	1.65	1.64	1.61	1.60	1.59	1.58	1.57	.10	22
2.15	2.07	2.03	1.98	1.94	1.91	1.89	1.85	1.84	1.82	1.80	1.78	.05	
2.98	2.83	2.75	2.67	2.58	2.53	2.50	2.42	2.40	2.36	2.33	2.31	.01	
1.35	1.33	1.32	1.31	1.30	1.29	1.29	1.28	1.28	1.27	1.27	1.26	.25	
1.78	1.73	1.70	1.67	1.64	1.62	1.61	1.58	1.57	1.56	1.54	1.53	.10	24
2.11	2.03	1.98	1.94	1.89	1.86	1.84	1.80	1.79	1.77	1.75	1.73	.05	
2.89	2.74	2.66	2.58	2.49	2.44	2.40	2.33	2.31	2.27	2.24	2.21	.01	
1.34	1.32	1.31	1.30	1.29	1.28	1.28	1.26	1.26	1.26	1.25	1.25	.25	
1.76	1.71	1.68	1.65	1.61	1.59	1.58	1.55	1.54	1.53	1.51	1.50	.10	26
2.07	1.99	1.95	1.90	1.85	1.82	1.80	1.76	1.75	1.73	1.71	1.69	.05	
2.81	2.66	2.58	2.50	2.42	2.36	2.33	2.25	2.23	2.19	2.16	2.13	.01	
1.33	1.31	1.30	1.29	1.28	1.27	1.27	1.26	1.25	1.25	1.24	1.24	.25	
1.74	1.69	1.66	1.63	1.59	1.57	1.56	1.53	1.52	1.50	1.49	1.48	.10	28
2.04	1.96	1.91	1.87	1.82	1.79	1.77	1.73	1.71	1.69	1.67	1.65	.05	
2.75	2.60	2.52	2.44	2.35	2.30	2.26	2.19	2.17	2.13	2.09	2.06	.01	
1.32	1.30	1.29	1.28	1.27	1.26	1.26	1.25	1.24	1.24	1.23	1.23	.25	
1.72	1.67	1.64	1.61	1.57	1.55	1.54	1.51	1.50	1.48	1.47	1.46	.10	30
2.01	1.93	1.89	1.84	1.79	1.76	1.74	1.70	1.68	1.66	1.64	1.62	.05	
2.70	2.55	2.47	2.39	2.30	2.25	2.21	2.13	2.11	2.07	2.03	2.01	.01	
1.30	1.28	1.26	1.25	1.24	1.23	1.22	1.21	1.21	1.20	1.19	1.19	.25	
1.66	1.61	1.57	1.54	1.51	1.48	1.47	1.43	1.42	1.41	1.39	1.38	.10	40
1.92	1.84	1.79	1.74	1.69	1.66	1.64	1.59	1.58	1.55	1.53	1.51	.05	
2.52	2.37	2.29	2.20	2.11	2.06	2.02	1.94	1.92	1.87	1.83	1.80	.01	
1.27	1.25	1.24	1.22	1.21	1.20	1.19	1.17	1.17	1.16	1.15	1.15	.25	
1.60	1.54	1.51	1.48	1.44	1.41	1.40	1.36	1.35	1.33	1.31	1.29	.10	60
1.84	1.75	1.70	1.65	1.59	1.56	1.53	1.48	1.47	1.44	1.41	1.39	.05	
2.35	2.20	2.12	2.03	1.94	1.88	1.84	1.75	1.73	1.68	1.63	1.60	.01	
1.24	1.22	1.21	1.19	1.18	1.17	1.16	1.14	1.13	1.12	1.11	1.10	.25	
1.55	1.48	1.45	1.41	1.37	1.34	1.32	1.27	1.26	1.24	1.21	1.19	.10	120
1.75	1.66	1.61	1.55	1.50	1.46	1.43	1.37	1.35	1.32	1.28	1.25	.05	
2.19	2.03	1.95	1.86	1.76	1.70	1.66	1.56	1.53	1.48	1.42	1.38	.01	
1.23	1.21	1.20	1.18	1.16	1.14	1.12	1.11	1.10	1.09	1.08	1.06	.25	
1.52	1.46	1.42	1.38	1.34	1.31	1.28	1.24	1.22	1.20	1.17	1.14	.10	200
1.72	1.62	1.57	1.52	1.46	1.41	1.39	1.32	1.29	1.26	1.22	1.19	.05	
2.13	1.97	1.89	1.79	1.69	1.63	1.58	1.48	1.44	1.39	1.33	1.28	.01	
1.22	1.19	1.18	1.16	1.14	1.13	1.12	1.09	1.08	1.07	1.04	1.00	.25	
1.49	1.42	1.38	1.34	1.30	1.26	1.24	1.18	1.17	1.13	1.08	1.00	.10	∞
1.67	1.57	1.52	1.46	1.39	1.35	1.32	1.24	1.22	1.17	1.11	1.00	.05	
2.04	1.88	1.79	1.70	1.59	1.52	1.47	1.36	1.32	1.25	1.15	1.00	.01	

TABLE D.4
Upper Percentage Points of the χ^2 Distribution

Example

$\Pr(\chi^2 > 10.85) = 0.95$
$\Pr(\chi^2 > 23.83) = 0.25$ for df = 20
$\Pr(\chi^2 > 31.41) = 0.05$



Degrees of freedom \ Pr	.995	.990	.975	.950	.900
1	392704×10^{-10}	157088×10^{-9}	982069×10^{-9}	393214×10^{-8}	.0157908
2	.0100251	.0201007	.0506356	.102587	.210720
3	.0717212	.114832	.215795	.351846	.584375
4	.206990	.297110	.484419	.710721	1.063623
5	.411740	.554300	.831211	1.145476	1.61031
6	.675727	.872085	1.237347	1.63539	2.20413
7	.989265	1.239043	1.68987	2.16735	2.83311
8	1.344419	1.646482	2.17973	2.73264	3.48954
9	1.734926	2.087912	2.70039	3.32511	4.16816
10	2.15585	2.55821	3.24697	3.94030	4.86518
11	2.60321	3.05347	3.81575	4.57481	5.57779
12	3.07382	3.57056	4.40379	5.22603	6.30380
13	3.56503	4.10691	5.00874	5.89186	7.04150
14	4.07468	4.66043	5.62872	6.57063	7.78953
15	4.60094	5.22935	6.26214	7.26094	8.54675
16	5.14224	5.81221	6.90766	7.96164	9.31223
17	5.69724	6.40776	7.56418	8.67176	10.0852
18	6.26481	7.01491	8.23075	9.39046	10.8649
19	6.84398	7.63273	8.90655	10.1170	11.6509
20	7.43386	8.26040	9.59083	10.8508	12.4426
21	8.03366	8.89720	10.28293	11.5913	13.2396
22	8.64272	9.54249	10.9823	12.3380	14.0415
23	9.26042	10.19567	11.6885	13.0905	14.8479
24	9.88623	10.8564	12.4011	13.8484	15.6587
25	10.5197	11.5240	13.1197	14.6114	16.4734
26	11.1603	12.1981	13.8439	15.3791	17.2919
27	11.8076	12.8786	14.5733	16.1513	18.1138
28	12.4613	13.5648	15.3079	16.9279	18.9392
29	13.1211	14.2565	16.0471	17.7083	19.7677
30	13.7867	14.9535	16.7908	18.4926	20.5992
40	20.7065	22.1643	24.4331	26.5093	29.0505
50	27.9907	29.7067	32.3574	34.7642	37.6886
60	35.5346	37.4848	40.4817	43.1879	46.4589
70	43.2752	45.4418	48.7576	51.7393	55.3290
80	51.1720	53.5400	57.1532	60.3915	64.2778
90	59.1963	61.7541	65.6466	69.1260	73.2912
100*	67.3276	70.0648	74.2219	77.9295	82.3581

*For df greater than 100 the expression $\sqrt{2\chi^2} - \sqrt{(2k-1)} = Z$ follows the standardized normal distribution, where k represents the degrees of freedom.

χ^2-table continued

.750	.500	.250	.100	.050	.025	.010	.005
.1015308	.454937	1.32330	2.70554	3.84146	5.02389	6.63490	7.87944
.575364	1.38629	2.77259	4.60517	5.99147	7.37776	9.21034	10.5966
1.212534	2.36597	4.10835	6.25139	7.81473	9.34840	11.3449	12.8381
1.92255	3.35670	5.38527	7.77944	9.48773	11.1433	13.2767	14.8602
2.67460	4.35146	6.62568	9.23635	11.0705	12.8325	15.0863	16.7496
3.45460	5.34812	7.84080	10.6446	12.5916	14.4494	16.8119	18.5476
4.25485	6.34581	9.03715	12.0170	14.0671	16.0128	18.4753	20.2777
5.07064	7.34412	10.2188	13.3616	15.5073	17.5346	20.0902	21.9550
5.89883	8.34283	11.3887	14.6837	16.9190	19.0228	21.6660	23.5893
6.73720	9.34182	12.5489	15.9871	18.3070	20.4831	23.2093	25.1882
7.58412	10.3410	13.7007	17.2750	19.6751	21.9200	24.7250	26.7569
8.43842	11.3403	14.8454	18.5494	21.0261	23.3367	26.2170	28.2995
9.29906	12.3398	15.9839	19.8119	22.3621	24.7356	27.6883	29.8194
10.1653	13.3393	17.1170	21.0642	23.6848	26.1190	29.1413	31.3193
11.0365	14.3389	18.2451	22.3072	24.9958	27.4884	30.5779	32.8013
11.9122	15.3385	19.3688	23.5418	26.2962	28.8454	31.9999	34.2672
12.7919	16.3381	20.4887	24.7690	27.5871	30.1910	33.4087	35.7185
13.6753	17.3379	21.6049	25.9894	28.8693	31.5264	34.8053	37.1564
14.5620	18.3376	22.7178	27.2036	30.1435	32.8523	36.1908	38.5822
15.4518	19.3374	23.8277	28.4120	31.4104	34.1696	37.5662	39.9968
16.3444	20.3372	24.9348	29.6151	32.6705	35.4789	38.9321	41.4010
17.2396	21.3370	26.0393	30.8133	33.9244	36.7807	40.2894	42.7956
18.1373	22.3369	27.1413	32.0069	35.1725	38.0757	41.6384	44.1813
19.0372	23.3367	28.2412	33.1963	36.4151	39.3641	42.9798	45.5585
19.9393	24.3366	29.3389	34.3816	37.6525	40.6465	44.3141	46.9278
20.8434	25.3364	30.4345	35.5631	38.8852	41.9232	45.6417	48.2899
21.7494	26.3363	31.5284	36.7412	40.1133	43.1944	46.9630	49.6449
22.6572	27.3363	32.6205	37.9159	41.3372	44.4607	48.2782	50.9933
23.5666	28.3362	33.7109	39.0875	42.5569	45.7222	49.5879	52.3356
24.4776	29.3360	34.7998	40.2560	43.7729	46.9792	50.8922	53.6720
33.6603	39.3354	45.6160	51.8050	55.7585	59.3417	63.6907	66.7659
42.9421	49.3349	56.3336	63.1671	67.5048	71.4202	76.1539	79.4900
52.2938	59.3347	66.9814	74.3970	79.0819	83.2976	88.3794	91.9517
61.6983	69.3344	77.5766	85.5271	90.5312	95.0231	100.425	104.215
71.1445	79.3343	88.1303	96.5782	101.879	106.629	112.329	116.321
80.6247	89.3342	98.6499	107.565	113.145	118.136	124.116	128.299
90.1332	99.3341	109.141	118.498	124.342	129.561	135.807	140.169

Source: Abridged from E. S. Pearson and H. O. Hartley, eds., *Biometrika Tables for Statisticians*, vol. 1, 3d ed., table 8, Cambridge University Press, New York, 1966. Reproduced by permission of the editors and trustees of *Biometrika*.

Department of Statistics

Stockholms universitet

Correction sheet

Date: 29/11-16

Room: Uggleviksssalen

Exam: Regressionsanalys/Ekonometri I

Course: Ekonometri

Anonymous code: EKR-0013

☒ I authorise the anonymous posting of my exam, in whole or in part, on the department homepage as a sample student answer.

NOTE! ALSO WRITE ON THE BACK OF THE ANSWER SHEET

Mark answered questions

1	2	3	4	5	6	7	8	9	Total number of pages
X	X	X	X						3
Teacher's notes 30	25	18	23						

Points	Grade	Teacher's sign.
96	A	PG

SU, DEPARTMENT OF STATISTICS

Room: UGGLEVIKSSALEN Anonymous code: EKR-0013 Sheet number: 1

EXERCISE 1

$n = 24$

$$\hat{Y}_t = \underset{\hat\beta_1}{1{,}31} - \underset{\hat\beta_2}{0{,}73} D_{2t} - \underset{\hat\beta_3}{0{,}77} D_{3t} - \underset{\hat\beta_4}{0{,}03} D_{4t} + \underset{\hat\beta_5}{2{,}34} X_t$$

Quarter j, $j = 2,3,4$ $(K = 5)$

$RSS = 7{,}94$
$ESS = 15{,}35$ $\Big\}$ $TSS = 23{,}29$

a) $H_0: \beta_2 = \beta_3 = \beta_4 = \beta_5 = 0$
H_1: at least one of $\beta_2, \beta_3, \beta_4, \beta_5 \neq 0$

test statistic: $F = \dfrac{ESS/(K-1)}{RSS/(n-K)} \overset{H_0}{\sim} F_{K-1;\, n-K}$

$$F_{obs} = \frac{15{,}35/4}{7{,}94/(\underbrace{24-5}_{19})} = 9{,}183$$

$F_{(0,05);4,19} = 2{,}90$

$\Rightarrow$ Since $F_{obs} = 9{,}183 > F_{(0,05);4,19} = 2{,}90$ we Reject H_0 and this means that at least one of the explanatory variables is statistically significant

OK

b) $$\hat{Y}_t = \underset{\hat\beta_1'}{1{,}16} - \underset{\hat\beta_2'}{0{,}27} D_{2t} - \underset{\hat\beta_3'}{0{,}39} D_{3t} + \underset{\hat\beta_4'}{0{,}01} D_{4t} + \underset{\hat\beta_5'}{2{,}03} X_t + \underset{\hat\beta_6'}{0{,}76} D_t X_t$$

$RSS = 6{,}61$ $(K = 6)$
$TSS = 23{,}29 \Rightarrow ESS = 16{,}68$
$\alpha = 0{,}05$

We define the model in (a) as the "old" model and the one in (b) as the "new" model and use the following test statistic for the "incremental" test:

$$F = \frac{(ESS_{new} - ESS_{old})/1}{RSS_{new}/(n-6)} \overset{H_0}{\sim} F_{1;18}$$

$$F_{obs} = \frac{(16{,}68 - 15{,}35)}{6{,}61/18} = 3{,}62$$

$F_{(0,05);1,18} = 4{,}41$

Ok

$\Rightarrow$ Since $F_{obs} = 3{,}62 < F_{(0,05);1,18}$, the null hypothesis is not rejected, so we can't conclude that the second model is significantly better than the first one using this test.

c) ~~$\hat{Y}_t = 0{,}03(-D_{1t} - D_{2t} - D_{3t} - D_{4t}) +$~~

$$\hat{Y}_t = -0{,}03(-D_{1t} - D_{2t} - D_{3t} - D_{4t}) + 1{,}31 - 0{,}73D_{2t} - 0{,}77D_{3t}$$

$$- 0{,}03D_{4t} + 2{,}34X_t$$

$\Leftrightarrow$

$$\hat{Y}_t = 1{,}28 + 0{,}03D_{1t} - 0{,}7D_{2t} - 0{,}74D_{3t} + 2{,}34X_t$$

If all the dummy variables are put to zero $\hat{Y}_t$ will give us the average value of Y in the 1st quarter given a certain value of X. This because the intercept is now interpretated as the average Y value in quarter 4, while the coefficients of the other dummies give us the difference ~~from~~ of the corresponding quarter from the 1st, in term of average Y value.

d) If we know X_t should be in the model but for some reasons we drop this variable the consequence will be that we will have an underfitting model and this would probably induce autocorrelation or serial correlation among the error terms:
$\Rightarrow cov(u_i; u_j) \neq 0 \quad (i \neq j)$

/30 OK

EXERCISE ②

$HRSA_i = \beta_1 + \beta_2 S_i + \beta_3 C_i + u_i$

$HRSA_i$ = # infections per thousand patients in hospital i
S_i = expenditure per patient on screening
C_i = " " on cleaning

Using regional data:

$$\overline{HRSA}_j = \beta_1 + \beta_2 \bar{S}_j + \beta_3 \bar{C}_j + \bar{u}_j$$

average of hospitals in region j

n_j hospitals in region j

a) We can rewrite the second model as follows:

$$\frac{\sum_{i=1}^{n_j} HRSA_i}{n_j} = \beta_1 + \beta_2 \frac{\sum_{i=1}^{n_j} S_i}{n_j} + \beta_3 \frac{\sum_{i=1}^{n_j} C_i}{n_j} + \frac{\sum_{i=1}^{n_j} u_i}{n_j}$$

As a consequence of this:

$$Var(\bar{u}_j) = Var\left(\frac{\sum u_i}{n_j}\right) = \frac{\sum_{i=1}^{n_j} Var(u_i)}{n_j^2} = \frac{n_j \sigma^2}{n_j^2} = \frac{\sigma^2}{n_j}$$

→ u_i's are assumed to be uncorrelated

So we can see that the variance of the error terms depends on n_j, which is diff in each region, thus the error terms do not have constant variance, there is heteroscedasticity

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If we use OLS estimators for the second model we will end up with a heteroscedastic model, this because the OLS estimates do not take in count that different regions have different variances. On the contrary, using WLS we would end up with an homoscedastic model.

OK

b) If we know the value of n_j for each region:

$$\sqrt{n_j}\, MRSA_j = \sqrt{n_j}\,\beta_1 + \sqrt{n_j}\,\beta_2\, S_j + \beta_3\sqrt{n_j}\, C_j + \underbrace{\sqrt{n_j}\,\bar{u}_j}$$

will be an homoscedastic model: $\bar{u}_j^*$

$$Var(\bar{u}_j^*) = Var(\sqrt{n_j}\,\bar{u}_j) = n_j \cdot Var(\bar{u}_j) = n_j \cdot \frac{\sigma^2}{n_j} = \sigma^2$$

So in the last model the variance of the error terms is constant.

OK

c) The problem with heteroscedasticity is that the OLS estimators are not efficient anymore and the consequence is that the estimates of the* ~~S.E. will be overestimated and thus the t-tests underestimated leading to unreliable t-test. A solution for this problem is to use White's robust estimates of the S.E.~~

* S.E.'s won't be correct leading thus to unreliable t-test.

A solution for this problem is to use White's robust estimations of standard errors. ~~these are usually higher in case of heteroscedasticity~~

Limitation of the robust method?

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EXERCISE ③

1)

(1) $\widehat{\log(Y_i)} = \hat{\beta}_1 + \hat{\beta}_2 \log(X_i)$

(2) $\widehat{\log(Y_i/X_i)} = \hat{\alpha}_1 + \hat{\alpha}_2 \log(X_i)$

From the second

$\log(Y_i) - \log(X_i) = \hat{\alpha}_1 + \hat{\alpha}_2 \log(X_i)$

$\Leftrightarrow$

$\log(Y_i) = \hat{\alpha}_1 + (\hat{\alpha}_2 + 1) \cdot \log(X_i)$

So we have

$\hat{\beta}_1 = \hat{\alpha}_1 = \bar{Y} - \hat{\beta}_2 \bar{X}$

$\hat{\beta}_2 = \hat{\alpha}_2 + 1 \Rightarrow \hat{\alpha}_2 = \hat{\beta}_2 - 1$ OK

* proof on sheet number 3

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EXERCISE ④

(a) F, the Runs test is a non parametric test, thus there is no need of the normal assumption of the disturbance variables in order to implement it. On the contrary the Durbin-Watson test relies on the assumption of normality of the error terms. OK

(b) F, when autocorrelation is present OLS estimators are still unbiased but they are inefficient OK

(c) T

If ~~...~~ the following is the "true" model:

$Y_i = \beta_1 + \beta_2 X_{2i} + \beta_3 X_{3t} + u_t$

but we use instead:

$Y_i = \beta_1' + \beta_2' X_{2i} + u_i'$

We will have that: $u_i' = \beta_3 X_{3t} + u_t$
so the error terms will be correlated because they all have X_{3i} in common. OK

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(d) F, because: $p_i = \frac{e^{\beta_1+\beta_2 X_{2i}}}{1+e^{\beta_1+\beta_2 X_{2i}}}$ / $\ln\left(\frac{p_i}{1-p_i}\right) = \beta_1+\beta_2 X_{2i}$

so p_i is not linear, while $\ln(\frac{p_i}{1-p_i})$ is linear

OK

(e) T, If we multiply Y with 5 we will have that

- $\hat{\beta}_2' = \frac{\sum(x_i-\bar{x})(5Y_i-5\bar{Y})}{\sum(x_i-\bar{x})^2} = 5\frac{\sum(x_i-\bar{x})(Y_i-\bar{Y})}{\sum(x_i-\bar{x})^2} = 5\hat{\beta}_2$

and

- $\hat{\beta}_1' = 5\bar{Y} + \hat{\beta}_2'\bar{X} = 5\bar{Y} + 5\hat{\beta}_2\bar{X} = 5(\bar{Y}+\hat{\beta}_2\bar{X}) = 5\hat{\beta}_1$

Where $\hat{\beta}_2', \hat{\beta}_1'$ are the estimates of the parameters in the model where Y is multiplied with 5 and $\hat{\beta}_2, \hat{\beta}_1$ are the estimates of the parameters in the first model

OK

(f) F,

$R^2 = 1 - \frac{RSS}{TSS}$

$\bar{R}^2 = 1 - \frac{RSS/(n-2)}{TSS/(n-1)} \neq R^2$

(*)(*) next page

~~g) If there is a measurement error we can rewrite the model as:~~

~~$Y_i + \varepsilon_i = \beta_1 + \beta_2 X_i + u_i \Rightarrow Y_i = \beta_1 + \beta_2 X_i + u_i - \varepsilon_i$~~

~~And would possibly lead to autocorrelation, but not to $cov(X_i, u_i) \neq 0$. So is FALSE~~

OK

(*) EXERCISE (2) continued:

$\hat{\beta}_2 = \frac{\sum(\ln(x_i)-\ln(\bar{x}))(\ln(Y_i)-\ln(\bar{Y}))}{\sum(\ln(x_i)-\ln(\bar{x}))^2}$

$\hat{\alpha}_2 = \frac{\sum(\ln(x_i)-\ln(\bar{x}))(\ln(Y_i/x_i)-\ln(\bar{Y}/\bar{x}))}{\sum(\ln(x_i)-\ln(\bar{x}))^2} =$

$= \frac{\sum(\ln(x_i)-\ln(\bar{x}))(\ln(Y_i)-\ln(x_i)-\ln(\bar{Y})+\ln(\bar{x}))}{\sum(\ln(x_i)-\ln(\bar{x}))^2} =$

$= \frac{\sum(\ln(x_i)-\ln(\bar{x}))\left[(\ln(Y_i)-\ln(\bar{Y}))-(\ln(x_i)-\ln(\bar{x}))\right]}{\sum(\ln(x_i)-\ln(\bar{x}))^2} =$

$\cancel{\sum \ln(X_i)}$

$$= \frac{\sum\left[(\ln(X_i) - \ln(\bar{X}))\cdot(\ln(Y_i) - \ln(\bar{Y}))\right] - \sum(\ln(X_i) - \ln(\bar{X}))^2}{\sum(\ln(X_i) - \ln(\bar{X}))^2} =$$

$$= \frac{\sum(\ln(X_i) - \ln(\bar{X}))\cdot(\ln(Y_i) - \ln(\bar{Y}))}{\sum(\ln(X_i) - \ln(\bar{X}))^2} - 1$$

$$\Rightarrow \hat{\alpha}_2 = \hat{\beta}_2 - 1 \qquad (\hat{\beta}_2 = \hat{\alpha}_2 + 1)$$

As a consequence:

$\cancel{\hat{\alpha}_1 = \ln(Y)}$

$$\hat{\alpha}_1 = \ln(\bar{Y}/\bar{X}) - \hat{\alpha}_2 \ln(\bar{X}) =$$

$$= \ln(\bar{Y}) - \cancel{\ln(\bar{X})} - \hat{\beta}_2 \ln(\bar{X}) + \cancel{\ln(\bar{X})} =$$

$$= \ln(\bar{Y}) - \hat{\beta}_2 \ln(\bar{X})$$

$$\Rightarrow \hat{\alpha}_1 = \hat{\beta}_1$$

(g) F,

$$Y_i = \beta_1 + \beta_2 X_i + u_i$$

Measurement error in X_i: if X_i^* is the true value, we will have:

$$Y_i = \beta_1 + \beta_2 X_i^* + \varepsilon_i + u_i$$

This could lead to autocorrelation among error terms.

OK

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